



STA3702

May/June 2017

STATISTICAL INFERENCE III

Duration 2 Hours 100 Marks

EXAMINATION PANEL AS APPOINTED BY THE DEPARTMENT

Use of a non-programmable pocket calculator is permissible

Closed book examination

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This paper consists of 4 pages

INSTRUCTIONS

- 1 Answer all questions
- 2 Show intermediate steps

Abbreviations	
MLE	Maximum Likelihood Estimator
MSE	Mean Square Error
pdf	probability density function
cdf	cumulative distribution function
MVUE	Minimum Variance Unbiased Estimator
MME	Method of Moments Estimator
CRLB	Cramer Rao Lower Bound
pmf	probability mass function
mgf	moment generating function
re	relative efficiency

TURN OVER

Question 1 [30]

A distribution belongs to the regular 1-parameter exponential family if among other regularity conditions its pdf or pmf has the form

$$f(x|\theta) = a(\theta)g(x) \exp\{b(\theta)r(x)\}, \ x \in \chi \subset (-\infty, \infty) \text{ and } \theta \in \Theta \subset (-\infty, \infty)$$

where g(x) > 0, r(x) is a function of x which does not depend on θ , and $a(\theta) > 0$ and $b(\theta) \neq 0$ are real valued functions of θ Furthermore, for this distribution (assuming the first and second derivatives of $a(\theta)$ and $b(\theta)$ exist)

$$E[r(X)] = -\frac{a'(\theta)}{a(\theta)b'(\theta)}$$

- (a) Write down the complete sufficient statistic for θ in terms of a random sample X_1, X_2, \dots, X_n from the distribution. Justify your answer. (1+1)
- (b) Write down a function of the complete sufficient statistic for θ (in terms of a random sample $X_1, X_2, ..., X_n$ from the distribution) that is a minimum variance unbiased estimator (MVUE) of E[r(X)] Justify your answer. (2+1)
- (c) Suppose that X is random sample (of size n = 1) from a distribution with pmf

$$p(x|\theta) = \begin{cases} \binom{m}{x} \theta^x (1-\theta)^{m-x} & \text{if } 0 < \theta < 1 \text{ and } x = 0, 1, 2, \dots, m, \\ 0 & \text{elsewhere} \end{cases}$$

(i) Show that $p(x|\theta)$ belongs to the regular 1-parameter exponential family by showing that the pmf can be expressed as

$$p(x|\theta) = a(\theta)g(x) \exp\{b(\theta)r(x)\}, \ x \in \chi \subset (-\infty, \infty) \text{ and } \theta \in \Theta \subset (-\infty, \infty)$$

(6)

(5)

- (ii) What is the complete sufficient statistic for θ^{γ} Justify your answer. (3)
- (iii) What is the mean of the complete sufficient statistic for θ in part (ii) above? Justify your answer.
- (iv) What is the maximum likelihood estimator (MLE) of θ^{γ} Justify your answer. (6)
- (v) What is the method of moments estimator (MME) of θ^{γ} Justify your answer. (2)
- (vi) What is the minimum variance unbiased estimator (MVUE) of θ° Justify your answer.

(3)

Question 2 [30]

Let X_1, X_2, \dots, X_n be a random sample from a distribution with pdf

$$f(x|\theta) = \begin{cases} \frac{\theta}{(1+x)^{1+\theta}} & \text{if } x > 0 \text{ and } \theta > 0, \\ 0 & \text{otherwise} \end{cases}$$

Given $E[\ln(1+X)] = \frac{1}{\theta}$

(a) Prove that
$$\prod_{i=1}^{n} (1 + X_i)$$
 is a minimal sufficient statistic for θ

(b) Suppose that
$$n = 200$$
, $\sum_{i=1}^{200} \ln(1+x_i) = 20$, $\sum_{i=1}^{200} 2 \ln(1+x_i) = 40$, and $\sum_{i=1}^{200} 3 \ln(1+x_i) = 60$ Determine

(1) the method of moments estimate
$$(MME)$$
 of θ , (3)

(ii) the maximum likelihood estimate
$$(MLE)$$
 of θ , (9)

(iii) the observed information,
$$I(\mathbf{x})$$
, of θ , and (4)

(iv) and the approximate standard errors of the estimates of
$$\theta$$
 in (i) and (ii) (6)

Let X_1, X_2, \dots, X_n be a random sample from the inverse Gaussian distribution with pdf

$$f(x|\theta_1, \theta_2) = \begin{cases} \left(\frac{\theta_2}{2\pi}\right)^{1/2} x^{-3/2} \exp\left\{-\frac{1}{2}\theta_2 \left(\frac{\sqrt{x}}{\theta_1} - \frac{1}{\sqrt{x}}\right)^2\right\} & \text{if } x > 0, \theta_1 > 0 \text{ and } \theta_2 > 0, \\ 0 & \text{otherwise} \end{cases}$$

- (a) Find the likelihood function and the log-likelihood function of θ_1 and θ_2
- (b) Show that the respective maximum likelihood estimators (MLE's) of θ_1 and θ_2 are

$$\hat{\theta}_1 = \bar{X} \text{ and } \hat{\theta}_2 = \frac{n}{\sum_{i=1}^n \left(\frac{\sqrt{x_i}}{\hat{\theta}_1} - \frac{1}{\sqrt{x_i}}\right)^2}$$

Hint: First find the MLE of θ_1 and then that of θ_2 . (12)

Question 4 [22]

A random variable X has probability density function

$$f(x|\theta) = \begin{cases} \theta e^{-\theta x} & \text{if } x > 0 \text{ and } \theta > 0, \\ 0 & \text{otherwise} \end{cases}$$

It is desired to test the null hypothesis H_0 $\theta=1$ against the alternative H_1 $\theta=2$ A random sample X_1 of size n=1 is to be used **Given** $\int f(x|\theta)dx=-e^{-\theta x}$

- (a) Calculate the level of significance and the power of the test which rejects H_0 if $X_1 \ge \frac{1}{2}$ (6+9)
- (b) Consider the test which rejects H_0 if $X_1 \ge c$ Find c for which the level of significance of the test is 0.1

 (7)

TOTAL [100]

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