

**STA3702**

October/November 2017

STATISTICAL INFERENCE III

Duration 2 Hours

100 Marks

EXAMINATION PANEL AS APPOINTED BY THE DEPARTMENT**Use of a non-programmable pocket calculator is permissible****Closed book examination****This examination question paper remains the property of the University of South Africa and may not be removed from the examination venue**

This paper consists of 4 pages

INSTRUCTIONS

- 1 Answer all questions
- 2 Show intermediate steps

Abbreviations	
<i>MLE</i>	Maximum Likelihood Estimator
<i>MSE</i>	Mean Square Error
<i>pdf</i>	probability density function
<i>cdf</i>	cumulative distribution function
<i>MVUE</i>	Minimum Variance Unbiased Estimator
<i>MME</i>	Method of Moments Estimator
<i>CRLB</i>	Cramer Rao Lower Bound
<i>pmf</i>	probability mass function
<i>mgf</i>	moment generating function
<i>re</i>	relative efficiency

[TURN OVER]

Question 1

[33]

A distribution belongs to the regular 1-parameter exponential family if among other regularity conditions its *pdf* or *pmf* has the form

$$f(y|\theta) = a(\theta)g(y) \exp\{b(\theta)r(y)\}, \quad y \in \mathcal{X} \subset (-\infty, \infty) \text{ and } \theta \in \Theta \subset (-\infty, \infty)$$

where $g(y) > 0$, $r(y)$ is a function of y which does not depend on θ , and $a(\theta) > 0$ and $b(\theta) \neq 0$ are real valued functions of θ . Furthermore for this distribution (assuming the first and second derivatives of $a(\theta)$ and $b(\theta)$ exist)

$$E[r(Y)] = -\frac{a'(\theta)}{a(\theta)b'(\theta)}$$

Suppose that Y_1, Y_2, \dots, Y_n is a random sample from this distribution

(a) Show that $f(y|\theta)$ can also be expressed as

$$f(y|\theta) = g(y) \exp\{b(\theta)r(y) - q(\theta)\}, \quad y \in \mathcal{X} \subset (-\infty, \infty) \text{ and } \theta \in \Theta \subset (-\infty, \infty)$$

(2)

(b) Prove that $\sum_{i=1}^n r(Y_i)$ is a minimal sufficient statistic for θ

(5)

(c) Justify that $\sum_{i=1}^n r(Y_i)$ is also a complete sufficient statistic for θ

(1)

(d) Write down a function of the complete sufficient statistic for θ that is a minimum variance unbiased estimator (*MVUE*) of $E[r(Y)]$. **Justify your answer**

(3)

(e) Suppose that X_1, X_2, \dots, X_n is random sample from a distribution with *pdf*

$$f(x|\theta) = \begin{cases} \frac{\theta}{(1+x)^{1+\theta}} & \text{if } x > 0 \text{ and } \theta > 0, \\ 0 & \text{otherwise} \end{cases}$$

(i) Show that $f(x|\theta)$ belongs to the regular 1-parameter exponential family by showing that the *pdf* can be expressed as

$$f(x|\theta) = a(\theta)g(x) \exp\{b(\theta)r(x)\}, \quad x > 0 \text{ and } \theta > 0$$

(5)

(ii) What is the mean of the complete sufficient statistic for θ ? **Justify your answer.**

(6)

(iv) What is the maximum likelihood estimator (*MLE*) of θ ? **Justify your answer.**

(6)

(v) What is the method of moments estimator (*MME*) of θ ? **Justify your answer.**

(2)

(vi) What is the minimum variance unbiased estimator (*MVUE*) of $\frac{1}{\theta}$? **Justify your answer**

(3)

[TURN OVER]

Question 2

[30]

Let X_1, X_2, \dots, X_n be a random sample from a distribution with *pdf*

$$f(x|\theta) = \begin{cases} \frac{1}{\theta} & \text{if } 0 < x \leq \theta, \\ 0 & \text{otherwise,} \end{cases}$$

where $\theta > 0$. Let $Y_n = \max\{X_1, X_2, \dots, X_n\}$. The probability density function of Y_n is

$$g(y|\theta) = \begin{cases} ny^{n-1}\theta^{-n} & \text{if } 0 < y \leq \theta, \\ 0 & \text{otherwise} \end{cases}$$

Suppose that the ten independent observations,

23, 23, 24, 23, 23, 23, 24, 22, 24, 25,

are from a population with probability density function $f(x|\theta)$

- (a) What are the method of moments estimator and estimate of θ ? (4)
- (b) What are the maximum likelihood estimator and estimate of θ ? (8)
- (c) Find the mean and the variance, in terms of θ and n , of the maximum likelihood estimator of θ . Is the estimator consistent? (10)
- (d) What is an estimate of the standard error of the maximum likelihood estimate of θ ? (3)
- (e) Find an approximate 95% confidence interval for θ . (5)

Question 3

[18]

Let X_1, X_2, \dots, X_n be a random sample from the distribution with *pdf*

$$f(x|\theta_1, \theta_2) = \begin{cases} \frac{\theta_1}{\theta_2^{\theta_1}} x^{\theta_1-1} & \text{if } 0 < x \leq \theta_2 \text{ and } \theta_1 > 0, \\ 0 & \text{otherwise} \end{cases}$$

- (a) Find the likelihood function and the log-likelihood function of θ_1 and θ_2 . (6)
- (b) Show that the respective maximum likelihood estimators (*MLE's*) of θ_1 and θ_2 are

$$\hat{\theta}_1 = \frac{n}{n \ln X_{(n)} - \sum_{i=1}^n \ln X_i} \text{ and } \hat{\theta}_2 = X_{(n)} = \max\{X_1, X_2, \dots, X_n\},$$

respectively. **Hint. First find the *MLE* of θ_2 and then that of θ_1** (12)

[TURN OVER]

Question 4

[19]

A random variable X has probability density function

$$f(x|\theta) = \frac{\exp\{x - \theta\}}{(1 + \exp\{x - \theta\})^2}, \quad -\infty < x < \infty \text{ and } -\infty < \theta < \infty,$$

and **cumulative** distribution function

$$F(x|\theta) = 1 - \frac{1}{1 + \exp\{x - \theta\}}, \quad -\infty < x < \infty \text{ and } -\infty < \theta < \infty$$

It is desired to test the null hypothesis $H_0: \theta = 0$ against the alternative $H_1: \theta = 1$ using one random observation X

(a) Calculate the level of significance and the power of the test which rejects H_0 if $X > 1$ (12)

(b) Consider the test which rejects H_0 if $X_1 \geq c$. Find c for which the level of significance of the test is 0.1 (7)

TOTAL [100]

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