## Tutorial letter 201/2/2018

Forecasting

## STA2604

Semester 2

## Department of Statistics

Solutions to Assignment 1

## QUESTION 1

(1.1) Forecasting helps in predicting future events and conditions based on present conditions. Here are two examples:

- Demography: Forecasting the population growth in a country for planning new infrastructures (e.g. roads, hospitals, schools).
- Business: forecasting the number of customers for a certain product in a shop for supply planning.
(1.2) The four components of a time series are the following:
- Trend: describes the upward or downward change for a time series during a given period of time.
- Seasonal variations: describes periodic patterns in a time series during a year and repeated in subsequent years.
- Cyclical variations: describes upward or downward changes around the trend that may occur after several years (e.g. five years).
- Irregular variations: describes changes in time series that follow no known (predictable) pattern.
(1.3) Let $y_{t}$ and $\widehat{y}_{t}$ denote the actual and predicted costs, respectively, in the table below:

| Weeks | Actual cost $y_{t}$ | Predicted cost $\hat{y}_{t}$ | $e_{t}$ | $\left\|e_{t}\right\|$ | $e_{t}^{2}$ | $A P E_{t}$ |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 75 | 73.3 | 1.7 | 1.7 | 2.89 | 2.2667 |
| 2 | 89.7 | 86.5 | 3.2 | 3.2 | 10.24 | 3.5674 |
| 3 | 123.6 | 129.4 | -5.8 | 5.8 | 33.64 | 4.6926 |
| 4 | 102 | 99.7 | 2.3 | 2.3 | 5.29 | 2.2549 |
| 5 | 139 | 142 | -3 | 3 | 9 | 2.1583 |
| Total |  |  |  | 16 | 61.06 | 14.9399 |

(1.4.1) The forecast error $e_{t}$ for each day is given in the fourth column of the table.
(1.4.2) MAD $=\frac{\sum_{t=1}^{5}\left|e_{t}\right|}{5}=\frac{16}{5}=3.2$.
(1.4.3) MSE $=\frac{\sum_{t=1}^{5} e_{t}^{2}}{5}=\frac{61.06}{5}=12.212$.
(1.4.4) MAPE $=\frac{\sum_{t=1}^{5} A P E_{t}}{5}=\frac{14.9399}{5}=2.9880$.

## QUESTION 2

(2.1) First construct a table with necessary information for all parts of the question.

| $t$ | $y_{t}$ | $t y_{t}$ | $t^{2}$ | $\widehat{y}_{t}$ | $e_{t}=y_{t}-\widehat{y}_{t}$ | $e_{t}^{2}$ | $d_{t}=\left(e_{t}-e_{t-1}\right)^{2}$ | $\left(y_{t}-\bar{y}_{t}\right)^{2}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 98 | 98 | 1 | 99.62 | -1.62 | 2.62 |  | 5662.56 |
| 2 | 100 | 200 | 4 | 109.44 | -9.44 | 89.03 | 61.11 | 5365.56 |
| 3 | 110 | 330 | 9 | 119.25 | -9.25 | 85.62 | 0.03 | 4000.56 |
| 4 | 150 | 600 | 16 | 129.07 | 20.93 | 438.03 | 910.98 | 540.56 |
| 5 | 140 | 700 | 25 | 138.89 | 1.11 | 1.24 | 392.74 | 1105.56 |
| 6 | 160 | 960 | 36 | 148.71 | 11.29 | 127.55 | 103.68 | 175.56 |
| 7 | 200 | 1400 | 49 | 158.52 | 41.48 | 1720.29 | 910.98 | 715.56 |
| 8 | 150 | 1200 | 64 | 168.34 | -18.34 | 336.40 | 3578.15 | 540.56 |
| 9 | 156 | 1404 | 81 | 178.16 | -22.16 | 491.01 | 14.57 | 297.56 |
| 10 | 170 | 1700 | 100 | 187.98 | -17.98 | 323.15 | 17.49 | 10.56 |
| 11 | 200 | 2200 | 121 | 197.79 | 2.21 | 4.87 | 407.33 | 715.56 |
| 12 | 190 | 2280 | 144 | 207.61 | -17.61 | 310.17 | 392.74 | 280.56 |
| 13 | 220 | 2860 | 169 | 217.43 | 2.57 | 6.61 | 407.33 | 2185.56 |
| 14 | 210 | 2940 | 196 | 227.25 | -17.25 | 297.45 | 392.74 | 1350.56 |
| 15 | 260 | 3900 | 225 | 237.06 | 22.94 | 526.04 | 1614.63 | 7525.56 |
| 16 | 258 | 4128 | 256 | 246.88 | 11.12 | 123.61 | 139.66 | 7182.56 |
| 136 | 2772 | 26900 | 1496 |  |  | 4883.69 | 9344.15 | 37655.00 |

The values of $\widehat{\beta}_{1}$ and $\widehat{\beta}_{0}$ are the following:

$$
\begin{gathered}
\widehat{\beta}_{1}=\frac{16 \sum_{t=1}^{16} t y_{t}-\sum_{t=1}^{16} t \sum_{t=1}^{16} y_{t}}{16 \sum_{t=1}^{16} t^{2}-\left(\sum_{t=1}^{16} t\right)^{2}}=\frac{16 \times 26900-136 \times 2772}{16 \times 1496-(136)^{2}}=\frac{53408}{5440}=9.8176 \\
\widehat{\beta}_{0}=\bar{y}_{t}-\widehat{\beta}_{1} \bar{t}=\frac{\sum_{t=1}^{16} y_{t}-\widehat{\beta}_{1} \sum_{t=1}^{16} t}{16}=\frac{2772-9.8176 \times 136}{16}=\frac{1436.8064}{16}=89.8004
\end{gathered}
$$

Thus, the fitted model is: $\widehat{y}_{t}=89.8004+9.8176 t$.
(2.2) The point forecast at month 21 is: $\widehat{y}_{17}=89.8004+9.8176 \times 17=256.6996$. If there was a seasonal index, this result should be multiplied by that index (or added that index). The computation of the $95 \%$ prediction interval needs some extra calculations. The fitted model and the values of $t$ in the first column were used to calculated the fitted values in the fifth column. The residuals $e_{t}$ in the sixth column are $e_{t}=y_{t}-\widehat{y}_{t}$. The squared residuals $e_{t}^{2}$ may now be used to calculate the mean square residual as: $M S E=\frac{e_{t}^{2}}{n-1}$ and thus its square root is $s=\sqrt{\frac{e_{t}^{2}}{n-2}}=\sqrt{\frac{4883.7}{14}}=18.68$ The $D$ distance is: $D=\frac{1}{n}+\frac{\left(t_{0}-\bar{t}\right)^{2}}{S S_{t t}}$ where $t_{0}=17, \bar{t}=\frac{136}{16}=8.5$,
and

$$
S S_{t t}=\sum_{t=1}^{16}(t-\bar{t})^{2}=\sum_{t=1}^{16} t^{2}-\frac{\left(\sum_{t=1}^{16} t\right)^{2}}{16}=1496-\frac{136^{2}}{16}=340
$$

Hence,

$$
D=\frac{1}{n}+\frac{\left(t_{0}-\bar{t}\right)^{2}}{S S_{t t}}=\frac{1}{16}+\frac{(17-8.5)^{2}}{340}=0.275
$$

The $95 \%$ prediction interval is: $\left[\widehat{y}_{t} \pm t_{[0.025]}^{(16-2)} s \sqrt{1+D}\right]$. That is: $\left[256.7 \pm t_{[0.025]}^{(14)} 18.68 \sqrt{1+0.275}\right]$ where $t_{[0.025]}^{(14)}=2.145$. Hence, the required $95 \%$ prediction interval is: [211.46; 301.94].
(2.3) The null and alternative hypotheses are:
$H_{0}$ : The error terms are not autocorrelated.
$H_{a}$ : The error terms are positively correlated.
The Durbin-Watson test statistic for positive autocorrelation is:

$$
d=\frac{\sum_{t=2}^{16}\left(e_{t}-e_{t-1}\right)^{2}}{\sum_{t=1}^{16} e_{t}^{2}}=\frac{9344.15}{4883.69}=1.91
$$

We will reject $H_{0}$ if $d<d_{L, 0.05}$, do not reject $H_{0}$ if $d>d_{L, 0.05}$ and fail to conclude $H_{0}$ if $d_{L, 0.05}<d<_{U, 0.05}$. For $n=16$, we have: $d_{L, 0.05}=1.10$ and $d_{U, 0.05}=1.37$. Since $d=1.91>_{U, 0.05}=1.37$, we do not reject $H_{0}$. We conclude that the error terms are not autocorrelated.
(2.4) The mean sales is: $\bar{y}_{t}=\frac{\sum_{t=1}^{16} y_{t}}{16}=\frac{2772}{16}=173.25$.

The total variation is: $\sum_{t=1}^{16}\left(y_{t}-\bar{y}_{t}\right)^{2}=37655$.
The total unexplained variation is: $\sum_{t=1}^{16} e_{t}^{2}=\sum_{t=1}^{16}\left(y_{t}-\widehat{y}_{t}\right)^{2}=4883.69$.
Thus, the coefficient of determination is:

$$
R^{2}=1-\frac{\sum_{t=1}^{16}\left(y_{t}-\widehat{y}_{t}\right)^{2}}{\sum_{t=1}^{16}\left(y_{t}-\bar{y}_{t}\right)^{2}}=1-\frac{4883.69}{37655}=0.8703 .
$$

The adjusted coefficient of determination is:

$$
R_{\mathrm{adj}}^{2}=\bar{R}^{2}=\left(R^{2}-\frac{k}{n-1}\right)\left(\frac{n-1}{n-(k+1)}\right)=\left(0.8703-\frac{1}{15}\right) \frac{15}{14}=0.8610
$$

(2.5) There is no point of calculating a VIF here since there is only one predictor variable.

## QUESTION 3

The first important thing is to correctly capture that data. The dataset should look as follows:

| $x_{1}$ | $x_{2}$ | $x_{1} x_{2}$ | $y$ |
| :---: | :---: | :---: | :---: |
| 0 | -2 | 0 | 8 |
| 0 | -1 | 0 | 9 |
| 0 | 0 | 0 | 9.1 |
| 0 | 1 | 0 | 10.2 |
| 0 | 2 | 0 | 10.4 |
| 1 | -2 | -2 | 10 |
| 1 | -1 | -1 | 10.3 |
| 1 | 0 | 0 | 12.2 |
| 1 | 1 | 1 | 12.6 |
| 1 | 2 | 2 | 13.9 |

(3.1) Model (obtained using Excel).

| SUMMARY OUTPUT |  |
| :--- | :---: |
| Regression Statistics |  |
| Multiple R | 0.9876 |
| R Square | 0.9754 |
| Adjusted R Square | 0.9630 |
| Standard Error | 0.3490 |
| Observations | 10 |

Note that overall, the model fits the data since $R^{2}$ and $R_{a d j}^{2}$ are large. However, this information is often misleading. Residual analysis is a more suitable approach.
ANOVA

|  | df | SS | MS | F | Significance F |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Regression | 3 | 28.9300 | 9.6433 | 79.1518 | 0.00003 |
| Residual | 6 | 0.7310 | 0.1218 |  |  |
| Total | 9 | 29.6610 |  |  |  |

Here, the ANOVA table again confirms that overall, the model fits the data since $F$ is large ( $F=79.1518$ ) with $p=0.00003<0.05$. However, again this information is often misleading. Residual analysis is a more suitable approach. The table of parameter estimates is given below, and indicates that both $x_{1}$ (Bacteria type) and $x_{2}$ (Time), and their interaction ( $x_{1} x_{2}$ ) have positive significant effects on $y$ (growth) since all $p$-values are small ( $\alpha=0.05$ is often used when no significance level is indicated).

|  | Coefficients | Standard Error | t Stat | P -value | Lower 95\% | Upper 95\% |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
| Intercept | 9.3400 | 0.1561 | 59.8341 | 0.0000 | 8.9580 | 9.7220 |
| X Variable 1 $\left(x_{1}\right)$ | 2.4600 | 0.2208 | 11.1435 | 0.0000 | 1.9198 | 3.0002 |
| X Variable 2 $\left(x_{2}\right)$ | 0.6000 | 0.1104 | 5.4359 | 0.0016 | 0.3299 | 0.8701 |
| X Variable 3 $\left(x_{1} x_{2}\right)$ | 0.4100 | 0.1561 | 2.6266 | 0.0392 | 0.0280 | 0.7920 |

Hence the predictive model is:

$$
\widehat{y}=9.34+2.46 x_{1}+0.6 x_{2}+0.41 x_{1} x_{2} .
$$

(3.2) Here $x_{1}=0$ and $x_{2}=0$, and thus $\hat{y}=9.34$ which is not very different from the observed value 9.1. The residual is: $9.1-9.34=-0.24$.
(3.3) Residuals are calculated as $e=y-\widehat{y}$ where $\widehat{y}=9.34+2.46 x_{1}+0.6 x_{2}+0.41 x_{1} x_{2}$. Manual calculation in Excel (as requested in the question) gives the following table:

| Observation | $x_{1}$ | $x_{2}$ | $x_{1} x_{2}$ | $y$ | $\hat{y}$ | $e$ |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 0 | -2 | 0 | 8 | 8.14 | -0.14 |
| 2 | 0 | -1 | 0 | 9 | 8.74 | 0.26 |
| 3 | 0 | 0 | 0 | 9.1 | 9.34 | -0.24 |
| 4 | 0 | 1 | 0 | 10.2 | 9.94 | 0.26 |
| 5 | 0 | 2 | 0 | 10.4 | 10.54 | -0.14 |
| 6 | 1 | -2 | -2 | 10 | 9.78 | 0.22 |
| 7 | 1 | -1 | -1 | 10.3 | 10.79 | -0.49 |
| 8 | 1 | 0 | 0 | 12.2 | 11.8 | 0.4 |
| 9 | 1 | 1 | 1 | 12.6 | 12.81 | -0.21 |
| 10 | 1 | 2 | 2 | 13.9 | 13.82 | 0.08 |

If an option to calculate residuals was ticked in Excel, we could obtain the following output (correct answer but the aim of the question was to assess if you know how the residuals are computed):

## RESIDUAL OUTPUT

| Observation | Predicted Y | Residuals |
| :---: | :---: | :---: |
| 1 | 8.14 | -0.14 |
| 2 | 8.74 | 0.26 |
| 3 | 9.34 | -0.24 |
| 4 | 9.94 | 0.26 |
| 5 | 10.54 | -0.14 |
| 6 | 9.78 | 0.22 |
| 7 | 10.79 | -0.49 |
| 8 | 11.8 | 0.4 |
| 9 | 12.81 | -0.21 |
| 10 | 13.82 | 0.08 |

Plot of residuals versus fitted values:

Residuals versus fitted values


Clearly, the residuals are randomly scattered around the line $e=0$. Thus, the assumption made in part (3.1) for a linear model with interaction is supported by residual analysis.
(3.4) For $x_{1}=1$, the fitted model is:

$$
\begin{aligned}
\widehat{y} & =9.34+2.36+0.6 x_{2}+0.41 x_{2} \\
& =11.8+1.01 x_{2}
\end{aligned}
$$

The mean value for $x_{2}$ is $\bar{x}_{2}=\frac{-2-1+0+1+2-2-1+0+1+2}{10}=0$ and

$$
\begin{aligned}
S_{x_{2} x_{2}} & =\sum_{i=1}^{10}\left(x_{i 2}-\bar{x}_{2}\right)^{2} \\
& =\sum_{i=1}^{10} x_{i 2}^{2}-\frac{\left(\sum_{i=1}^{10} x_{i 2}\right)^{2}}{10} \\
& =(-2)^{2}+(-1)^{2}+0^{2}+1^{2}+2^{2}+(-2)^{2}+(-1)^{2}+0^{2}+1^{2}+2^{2}-0 \\
& =20 .
\end{aligned}
$$

The distance value at $x_{2}=1$ is thus:

$$
\begin{aligned}
D & =\frac{1}{n}+\frac{\left(x_{02}-\bar{x}_{2}\right)^{2}}{S_{x_{2} x_{2}}} \\
& =\frac{1}{10}+\frac{(1-0)^{2}}{20} \\
& =\frac{1}{10}+\frac{1}{20} \\
& =0.15 .
\end{aligned}
$$

The fitted value at $x_{2}=1$ is $\widehat{y}=11.8+1.01=12.81$ and for $\alpha=0.1$, we have $t_{[\alpha / 2]}^{(10-2)}=t_{[0.05]}^{(8)}=1.860$. Hence, the $90 \%$ confidence interval of the expected growth (mean value) for type $B$ bacteria at time $x_{2}=1$ is:

$$
\begin{aligned}
{\left[\widehat{y} \pm t_{[0.05]}^{(8)} s \sqrt{D}\right] } & =[12.81 \pm 1.860 \times 0.349 \sqrt{0.15}] \\
& =[12.81 \pm 0.25] \\
& =[12.56 ; 13.06] .
\end{aligned}
$$

(3.5) The $90 \%$ prediction interval of the expected growth (mean value) for type $B$ bacteria at time $x_{2}=1$ is:

$$
\begin{aligned}
{\left[\widehat{y} \pm t_{[0.05]}^{(8)} s \sqrt{1+D}\right] } & =[12.81 \pm 1.860 \times 0.349 \sqrt{1.15}] \\
& =[12.81 \pm 0.70] \\
& =[12.11 ; 13.51] .
\end{aligned}
$$

(3.6) The prediction interval is wider than the confidence interval, but this is expected to be so.

## QUESTION 4

(4.1) The dummy variables are defined in the following table:

| Quarter | $Q_{2}$ | $Q_{3}$ | $Q_{4}$ |
| :--- | :---: | :---: | :---: |
| 1 | 0 | 0 | 0 |
| 2 | 1 | 0 | 0 |
| 3 | 0 | 1 | 0 |
| 4 | 0 | 0 | 1 |

(4.2) The plot of $y_{t}$ versus $t$ (with $t=1,2, \ldots, 20$ ) is given below:

Plot of $y \_t$ versus $t$

(4.3) It appears to be a constant seasonal variation since there are similar quarter patterns throughout the five years. However, the trend increases.
(4.4) The output for the fitted model is the following:

| SUMMARY OUTPUT |  |
| :--- | :---: |
| Regression Statistics |  |
| Multiple R | 0.9889 |
| R Square | 0.9779 |
| Adjusted R Square | 0.9721 |
| Standard Error | 2.2449 |
| Observations | 20 |


| ANOVA |  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: |
|  | df | SS | MS | F | Significance F |
| Regression | 4 | 3350.9577 | 837.7394 | 166.2352 | 0.0000 |
| Residual | 15 | 75.5923 | 5.0395 |  |  |
| Total | 19 | 3426.5500 |  |  |  |


|  | Coefficients | Standard Error | t Stat | P-value | Lower 95\% | Upper 95\% |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
| Intercept | 11.3317 | 1.2889 | 8.7919 | 0.0000 | 8.5845 | 14.0789 |
| $t$ | 0.4965 | 0.0898 | 5.5281 | 0.0001 | 0.3051 | 0.6879 |
| $Q_{2}$ | 19.1035 | 1.4226 | 13.4284 | 0.0000 | 16.0713 | 22.1358 |
| $Q_{3}$ | 31.5063 | 1.4290 | 22.0484 | 0.0000 | 28.4606 | 34.5521 |
| $Q_{4}$ | 3.9106 | 1.4451 | 2.7060 | 0.0163 | 0.8304 | 6.9908 |

This table indicates that sales significantly increase with time and there is increase seasonal variations for quarters 2, 3, 4 compared to quarter 1 (all p-values very small). The prediction model is:

$$
\widehat{y}=11.3317+0.4965 t+19.1035 Q_{2}+31.5063 Q_{3}+3.9106 Q_{4} .
$$

(4.5) $t=17$ corresponds to the first quarter of year 4, that is $Q_{2}=Q_{3}=Q_{4}=0$, and thus the fitted model can be written as: $\widehat{y}=11.3317+0.4965 t$. Hence, the point forecast is
$\widehat{y}_{17}=11.3317+0.4965 \times 17=19.7722$. From the regression ANOVA table, $s=\sqrt{M S E}=\sqrt{5.0395}=2.2449$ (same as standard error), and since there were 5 parameters to estimate, we have

$$
t_{[\alpha / 2]}^{(n-p)}=t_{[0.025]}^{(20-5)}=t_{[0.025]}^{(15)}=2.131 .
$$

The mean value of $t$ is:

$$
\bar{t}=\frac{\sum_{t=1}^{20} t}{20}=\frac{1+2+\cdots+20}{20}=10.5 .
$$

Then $(17-\bar{t})^{2}=(17-10.5)^{2}=42.25$. Calculations for $\sum_{t=1}^{20}(t-\bar{t})^{2}$ can be read from the following table.

| Year | Quarter | t | $Q_{2}$ | $Q_{3}$ | $Q_{4}$ | $y_{t}$ | $t-\bar{t}$ | $(t-\bar{t})^{2}$ |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 1 | 1 | 0 | 0 | 0 | 12 | -9.5 | 90.25 |
| 1 | 2 | 2 | 1 | 0 | 0 | 31 | -8.5 | 72.25 |
| 1 | 3 | 3 | 0 | 1 | 0 | 42 | -7.5 | 56.25 |
| 1 | 4 | 4 | 0 | 0 | 1 | 21 | -6.5 | 42.25 |
| 2 | 1 | 5 | 0 | 0 | 0 | 13 | -5.5 | 30.25 |
| 2 | 2 | 6 | 1 | 0 | 0 | 33 | -4.5 | 20.25 |
| 2 | 3 | 7 | 0 | 1 | 0 | 45 | -3.5 | 12.25 |
| 2 | 4 | 8 | 0 | 0 | 1 | 19 | -2.5 | 6.25 |
| 3 | 1 | 9 | 0 | 0 | 0 | 15 | -1.5 | 2.25 |
| 3 | 2 | 10 | 1 | 0 | 0 | 36 | -0.5 | 0.25 |
| 3 | 3 | 11 | 0 | 1 | 0 | 52 | 0.5 | 0.25 |
| 3 | 4 | 12 | 0 | 0 | 1 | 21 | 1.5 | 2.25 |
| 4 | 1 | 13 | 0 | 0 | 0 | 15 | 2.5 | 6.25 |
| 4 | 2 | 14 | 1 | 0 | 0 | 37 | 3.5 | 12.25 |
| 4 | 3 | 15 | 0 | 1 | 0 | 52 | 4.5 | 20.25 |
| 4 | 4 | 16 | 0 | 0 | 1 | 22 | 5.5 | 30.25 |
| 5 | 1 | 17 | 0 | 0 | 0 | 24 | 6.5 | 42.25 |
| 5 | 2 | 18 | 1 | 0 | 0 | 40 | 7.5 | 56.25 |
| 5 | 3 | 18 | 0 | 1 | 0 | 50 | 7.5 | 56.25 |
| 5 | 4 | 20 | 0 | 0 | 1 | 23 | 9.5 | 90.25 |
| Total |  |  |  |  |  |  |  | 649 |

Hence, the $95 \%$ confidence interval (not prediction interval) of $y_{17}$ is thus:

$$
\left[\widehat{y}_{17} \pm s t_{[0.025]}^{(15)} \sqrt{\frac{1}{20}+\frac{(17-\bar{t})^{2}}{\sum_{t=1}^{20}(t-\bar{t})^{2}}}\right]
$$

that is:

$$
\left[19.7722 \pm 2.2449 \times 2.131 \sqrt{\frac{1}{20}+\frac{42.25}{649}}\right]=[19.7722 \pm 1.6230]=[18.1492 ; 21.3952] .
$$

