# **Tutorial letter 202/1/2017**

# Applied Statistics II STA2601

**Semester 1** 

**Department of Statistics** 

Solutions to Assignment 02





Define tomorrow.

## **QUESTION 1**

(a)

	$x_i$	$(x_i - \overline{x})$	$(x_i - \overline{x})^2$	$(x_i - \overline{x})^3$	$(x_i - \overline{x})^4$
	0.48	-0.13	0.0169	-0.002197	0.00028561
	0.96	0.35	0.1225	0.042875	0.01500625
	0.52	-0.09	0.0081	-0.000729	0.00006561
	0.36	-0.25	0.0625	-0.015625	0.00390625
	0.49	-0.12	0.0144	-0.001728	0.00020736
	÷	:	:		÷
	÷		:		:
	0.74	0.13	0.0169	0.002197	0.00028561
	0.55	-0.06	0.0036	-0.000216	0.00001296
	0.49	-0.12	0.0144	-0.001728	0.00020736
	0.58	-0.03	0.0009	-0.000027	0.0000081
	0.64	0.03	0.0009	0.000027	0.0000081
Total	36.6	0	1.4572	0.116088	0.104689

$$n = 60 \qquad \sum X = 36.6 \qquad \overline{X} = 0.61$$
  
$$\sum (X_i - \overline{X})^2 = 1.4572 \qquad \sum (X_i - \overline{X})^3 = 0.116088 \qquad \sum (X_i - \overline{X})^4 = 0.104689$$

## (b) (i) Test for skewness:

 $H_0: \ \ {\rm The \ distribution \ is \ normal} \ \ \left(\Rightarrow \beta_1=0\right).$ 

 $H_1: \ \beta_1 \neq 0.$ 

(Please note: The alternative must be two-sided. There is no indication of a one-sided test.)

(9)

The critical value is 0.492. Reject  $H_0$  if  $\beta_1 < -0.492$  or  $\beta_1 > 0.492$  or  $|\beta_1| > 0.492$ .

Now 
$$\beta_1 = \frac{\frac{1}{n} \Sigma (X_i - \overline{X})^3}{\left(\sqrt{\frac{1}{n} \Sigma (X_i - \overline{X})^2}\right)^3} = \frac{\frac{1}{60} (0.116088)}{\left(\sqrt{\frac{1}{60} (1.4572)}\right)^3}$$
$$= \frac{0.0019348}{\left(\sqrt{0.024286666}\right)^3}$$
$$= \frac{0.0019348}{(0.1558418)^3}$$
$$= \frac{0.0019348}{0.003784877}$$
$$\simeq 0.5112$$

Since 0.5112 > 0.492 we reject  $H_0$  at the 10% level of significance level and conclude that the distribution is not symmetric.

(7)

#### (ii) Test for kurtosis:

We have to test:

 $H_0$ : The distribution is normal ( $\Rightarrow \beta_2 = 3$ ).

 $H_1: \beta_2 \neq 3.$ 

Since we have a sample size greater than 50, the test is based on  $\beta_2$  (page 111 in the study guide).

The size of the sample, n = 60, thus 60 is between 50 and 75, we need to interpolate the critical values.

From table B (page 111 in the study guide):

The upper 5% percentage point for  $B_2$  is

$$3.99 + \frac{(60 - 50)}{75 - 60} (3.87 - 3.99) = 3.99 + \frac{10}{25} (-0.12) = 3.942.$$

The lower 5% percentage point for A is

$$2.15 + \frac{(60 - 50)}{75 - 60} (2.27 - 2.15) = 2.15 + \frac{10}{25} (0.12) = 2.198.$$

We reject  $H_0$  at the 10% significance level if  $B_2 < \text{lower 5\% point or } B_2 > \text{upper 5\% point}$  in table C.

The critical values are 3.942 and 2.198. Reject  $H_0$  if  $B_2 < 2.198$  or  $B_2 > 3.942$ .

Now the value of the test statistic is

$$\beta_2 = \frac{\frac{1}{n} \sum_{i=1}^n (X_i - \overline{X})^4}{\left[\frac{1}{n} \sum_{i=1}^n (X_i - \overline{X})^2\right]^2}$$
$$= \frac{\frac{1}{60} (0.104689)}{\left[\frac{1}{60} (1.4572)\right]^2}$$
$$= \frac{0.001744816}{[0.024286666]^2}$$
$$= \frac{0.001744816}{0.000589842}$$
$$\approx 2.9581$$

Since 2.198 < 2.9581 < 3.942, we do not reject  $H_0$  at the 10% level of significance and conclude that the distribution does have the kurtosis of a normal distribution.

- (7)
- (iii) No, the sample failed one test and we conclude that the distribution does not originate from a normal distribution.

(c) (i) If we divide the observations into 6 classes with equal expected frequencies, it means that  $\pi_i = \frac{1}{6}$  for each interval  $\Rightarrow n\pi_i = 10$ . The biggest problem is to *determine the interval limits* in terms of the *X*-scale such that each interval has a probability of  $\frac{1}{6} = 0.167$ .

We start with the standardised n(0; 1) scale (as always) and transform back to the *X*-scale by making use of

$$Z = \frac{X - \mu}{\sigma} = \frac{X - 0.6}{0.16}.$$



From the first interval we have  $P[Z \le -b] = 0.167$  and we find *b* by working with the positive mirror image. i.e.,  $P(Z \ge b) = 0.167 \implies P(Z \le b) = 0.833$  (this probability is 1 - 0.167).

From table II we get that  $\Phi(0.966) = 0.833 \Rightarrow b = 0.966$ .

The first interval is where  $Z \le -0.966$  [Note that  $P(Z \le -0.966) = 0.167$ .

$$\implies \frac{X - \mu}{\sigma} = \frac{X - 0.6}{0.16} \leq -0.966$$
$$X - 0.6 \leq -0.966 \times 0.16$$
$$X - 0.6 \leq -0.15456$$
$$X \leq 0.6 - 0.15456$$
$$X \leq 0.44544$$
$$X \leq 0.44544$$

Hence the first interval is where  $X \le 0.445$ .

The second interval is where  $-b \le Z \le -a$ . From the sketch above the value "*a*" is found from Table II as  $\Phi(0.432) = P(Z \le a) = 0.5 + 0.167 = 0.667$ . Thus -a = -0.432.

Thus

-0.966	$\leq$	Ζ	$\leq$	-0.432
-0.966	$\leq$	$\frac{X - 0.6}{0.16}$	$\leq$	-0.432
-0.966 ×0.16	$\leq$	X - 0.6	$\leq$	$-0.432 \times 0.16$
-0.15456	$\leq$	X - 0.6	$\leq$	-0.06912
0.6 - 0.15456	$\leq$	X	$\leq$	0.6 - 0.06912
0.44544	$\leq$	X	$\leq$	0.53088
0.445	<	X	<	0.531

The second interval is  $0.445 \le X \le 0.531$ .

(10)

(1)

(ii) Expected frequency is 
$$n\hat{\pi}_i = \frac{1}{6} \times 60 = 10$$
.

Table of observed and expected frequencies		
Class interval	$O_i$	$\hat{e}_i = n\hat{\pi}_i$
X < 0.445	10	10
$0.445 \le X < 0.531$	8	10
$0.531 \le X < 0.600$	14	10
$0.600 \le X < 0.669$	7	10
$0.669 \le X < 0.755$	12	10
$X \ge 0.755$	9	10
Totals	60	60

(iii) Goodness of fit test:

We have to test

 $H_0$ : The observations come from a  $n(0.6, 0.16^2)$  distribution

 $H_1$ : The observations do not come from a  $n(0.6, 0.16^2)$  distribution.

We use the test statistic

$$Y^{2} = \sum_{i=1}^{6} \frac{(N_{i} - n\pi_{i})^{2}}{n\pi_{i}} \quad (\text{page 98 study guide})$$
  

$$\therefore Y^{2} = \frac{1}{10} \left( (10 - 10)^{2} + (8 - 10)^{2} + (14 - 10)^{2} + (7 - 10)^{2} + (12 - 10)^{2} + (9 - 10)^{2} \right)$$
  

$$= \frac{1}{10} (0 + 4 + 16 + 9 + 4 + 1)$$
  

$$= \frac{34}{10}$$
  

$$= 3.4$$

The critical value is

$$\chi^{2}_{\alpha;k-r-1} = \chi^{2}_{0.05;6-0-1} \\ = \chi^{2}_{0.05;5} \\ = 11.0706$$

Reject  $H_0$  if  $Y^2 \ge 11.0706$ .

Since 3.4 < 11.0706, we cannot reject  $H_0$  at the 5% level of significance and conclude that the data follow a normal distribution.

FOR YOUR INFORMATION: For the degrees of freedom k - r - 1 where k is the number of classes and r is the number of estimated parameters.

Test	Value of r	Parameter unknown
$n(20, 5^2)$	r = 0	
$n(20,\sigma^2)$	r = 1	$\sigma$ unknown
$n(\mu, 5^2)$	r = 1	$\mu$ unknown
$n\left(\mu,\sigma^2\right)$	r = 2	both $\mu$ and $\sigma$ unknown

(10)

[45]

# **QUESTION 2**

- (a) Start the JMP program.
  - > Enter Job satisfaction in the first column and label it Job satisfaction.

(make sure to change the scale to nominal)

> Enter *Type of work* in the second column and label it *Type of work*.

(make sure to change the scale to ordinal)

> Enter the frequency in the third column and label it <u>Count</u>.

Your data should look like this.

Aggressiveness	Sales	Count
Satisfied	White collar	81
Satisfied	Blue collar	124
Satisfied	Menial	44
Dissatisfied	White collar	49
Dissatisfied	Blue collar	76
Dissatisfied	Menial	56

This is a chi-square test of association. To fit the model:

> Choose Analyze>Fit *Y* by *X* with <u>Job satisfaction</u> as *X*, Factor and <u>Type of work</u> as *Y*, Response and <u>Count</u> as Freq.

> <u>Click Ok</u>.



The **Mosaic Plot** shows that the sample sizes for satisfied and dissatisfied are almost in the ratio 3:2. The mosaic plot shows that the proportion of blue collar who are satisfied was almost twice those in white collar and three times those in menial. Thus the blue collar comprises almost half of those satisfied. The proportion of those dissatisfied who are blue collar was slightly higher than the other types of works who were almost the same. Thus, the proportions seem to be not the same as evidenced by horizontal lines which are not in alignment. The hypothesis of no association might be rejected.

(11)

(b)  $H_0$ : Job satisfaction is not related to type of work.

 $H_1$ : Job satisfaction is related to type of work.

(2)

(c) The test statistic is 
$$Y^2 = \sum_{k=1}^{k} \frac{(\text{Observed} - \text{Expected})^2}{\text{Expected}}$$
 and the value is  $Y^2 = 10.342$ .

(2)

(d) No, the row percentages seems to be different. For dissatisfied are 41.99%, 30.94%, 27.07% and for satisfied are 49.80%, 17.67% and 32.53% for blue collar, menial and white collar respectively. One might expect the null hypothesis to be rejected.

(3)

(e) The critical value is  $\chi^2_{0.05;2} = 5.99147$ . Since 10.342 > 5.99147, we reject  $H_0$  at the 5% level of significance and conclude that job satisfaction is related to type of work.

Alternatively the *p*-value is = 0.0057. Since 0.0057 < 0.05, we reject  $H_0$  at the 5% level of significance and conclude that job satisfaction is related to type of work.

(2)

[20]

# **QUESTION 3**

 $H_0$ : There is no difference in opinion between psychologists and psychiatrists regarding behavioural modification as a useful technique.

 $H_1$ : Psychiatrists feel less favourable towards the usefulness of behavioural modification

For this  $2 \times 2$  table for the exact test is

Is behavioural m	odifica	ation use	ful?	
	Yes No Row total			
Psychologists	5	2	7	
Psychiatrists	1	4(=x)	5	
Column total	6	6	12	
		1		
		п		

Now k = 5, n = 6 and x = 4 [Note: k = 5, n = 6 and x = 1 thus  $P(X \le 1) = 0.121$ ]

The alternative "Psychiatrists feel less favourable towards the usefulness of behavioural modification" would imply a large value of *x* to reject  $H_0$  i.e.,  $P(X \ge x) = \alpha$ .

Now x = 4 and

$$P(X \ge x) = 1 - P(X < x - 1)$$
  

$$P(X \ge 4) = 1 - P(X \le 3)$$
  

$$= 1 - 0.879$$
  

$$= 0.121 \text{ (from table D study guide p131)}$$

Since 0.121 > 0.05, we do not reject  $H_0$  at the 5% level of significance and conclude that there is no difference in opinion between psychologists and psychiatrists regarding behavioural modification as a useful technique.

[10]

# **QUESTION 4**

(a) If 
$$U = \frac{\Sigma (X_i - \mu)^2}{\sigma^2}$$
 then  $U \sim \chi_n^2$  (result 1.2).

Then

$$1 - \alpha = P\left(\chi_{1-\frac{1}{2}\alpha;n}^{2} < U < \chi_{\frac{1}{2}\alpha;n}^{2}\right)$$
$$= P\left[\chi_{1-\frac{1}{2}\alpha;n}^{2} < \frac{\Sigma(X_{i}-\mu)^{2}}{\sigma^{2}} < \chi_{\frac{1}{2}\alpha;n}^{2}\right]$$
$$= P\left[\frac{1}{\chi_{\frac{1}{2}\alpha;n}^{2}} < \frac{\sigma^{2}}{\Sigma(X_{i}-\mu)^{2}} < \frac{1}{\chi_{1-\frac{1}{2}\alpha;n}^{2}}\right]$$
$$= P\left[\frac{\Sigma(X_{i}-\mu)^{2}}{\chi_{\frac{1}{2}\alpha;n}^{2}} < \sigma^{2} < \frac{\Sigma(X_{i}-\mu)^{2}}{\chi_{1-\frac{1}{2}\alpha;n}^{2}}\right]$$

Thus the 100  $(1 - \alpha)$ % two-sided confidence interval for  $\sigma^2$  is given by  $\left[\frac{\Sigma (X_i - \mu)^2}{\chi^2_{\frac{1}{2}\alpha;n}}; \frac{\Sigma (X_i - \mu)^2}{\chi^2_{1-\frac{1}{2}\alpha;n}}\right]$ 

(b) (i) We have to test 
$$H_0: \sigma_1^2 = \sigma_2^2$$
  
against  $H_1: \sigma_1^2 \neq \sigma_2^2$   
 $n_1 = 17$   $S_1^2 = (1.12)^2 = 1.2544$   $n_2 = 12$   $S_2^2 = (2.36)^2 = 5.5696$ 

(7)

The test statistic is

$$F = \frac{\sigma_2^2}{\sigma_1^2} \times \frac{S_1^2}{S_2^2}$$
$$= 1 \times \frac{1.2544}{5.5696}$$
$$\approx 0.2252$$

The critical values is  $F_{\alpha/2;n_1-1;n_2-1} = F_{0.025;16;11} \approx 3.33$  (or actual is  $F_{0.025;16;11} = 3.30$  (internet F-calculator)) and  $F_{1-\alpha/2;n_1-1;n_2-1} = \frac{1}{F_{\alpha/2;n_2-1;n_1-1}} = \frac{1}{F_{0.025;11;16}} = \frac{1}{\frac{1}{\frac{1}{2}(2.99+2.89)}} = \frac{1}{\frac{1}{2.94}} \approx 0.3401$ . Reject  $H_0$  if F > 3.33 or F < 0.3401. Since 0.2252 < 0.3401, we reject  $H_0$  at the 5% level of significance and conclude that the variances are not equal i.e.  $\sigma_1^2 \neq \sigma_2^2$ .

(ii) The 95% confidence interval for  $\frac{\sigma_1^2}{\sigma_2^2}$  is  $P\left(F_{1-\frac{\alpha}{2};n_2-1;n_1-1} < \frac{\sigma_1^2}{\sigma_2^2}\frac{S_2^2}{S_1^2} < F_{\frac{\alpha}{2};n_2-1;n_1-1}\right) = 1 - \alpha$ (10)

$$\left[\frac{F_{1-\frac{\alpha}{2};n_2-1;n_1-1}}{S_2^2/S_1^2};\frac{F_{\frac{\alpha}{2};n_2-1;n_1-1}}{S_2^2/S_1^2}\right]$$

$$\alpha = 0.05, \alpha/2 = 0.025$$

$$F_{1-\frac{\alpha}{2};n_2-1;n_1-1} = F_{0.975;11;16} = \frac{1}{F_{0.025;16;11}} \approx \frac{1}{3.33} \approx 0.3003$$

$$F_{\frac{\alpha}{2};n_2-1;n_1-1} = F_{0.025;11;16} = \frac{1}{2} (2.99 + 2.89) = 2.94$$

$$\therefore \text{The 95\% confidence interval is}$$

$$\begin{bmatrix} \frac{F_{1-\frac{\alpha}{2};n_2-1;n_1-1}}{S_2^2/S_1^2}; \frac{F_{\frac{\alpha}{2};n_2-1;n_1-1}}{S_2^2/S_1^2} \end{bmatrix} \\ \begin{bmatrix} 0.3003\\ \overline{5.5696/1.2544}; \frac{2.94}{5.5696/1.2544} \end{bmatrix} \\ \begin{bmatrix} 0.3003\\ \overline{4.44005102}; \frac{2.94}{4.44005102} \end{bmatrix} \\ \begin{bmatrix} 0.0676; 0.6622 \end{bmatrix}. \end{bmatrix}$$

We are 95% confident that the confidence interval for  $\frac{\sigma_1^2}{\sigma_2^2}$  lies between 0.0676 to 0.6622. (8)

[25]

[100]