## Tutorial Letter 203/1/2014

## Applied Statistics II STA2601

## Semester 1

## Department of Statistics

## Solutions to Assignment 3

## QUESTION 1

(a)

(i) The assumptions are:

- observations are independent
- the data follows a normal distribution

Now based on the assumption of independent observations and the assumption that the IQ scores have a normal distribution (i.e., the sample comes from a normal population) we may assume that

$$
T=\frac{\sqrt{n}\left(\bar{X}-\mu_{0}\right)}{S} \sim t_{n-1}
$$

Are they met? The people were drawn randomly, thus the assumption of independent observations is met.

The normality assumption is not violated because from the JMP graphical output we see that the normal curve does fit the histogram very well. The box plot depicts a symmetric distribution and the points seem not to be deviating from the diagonal on the Normal Quantile Plot.
(ii) No.

We have to test $H_{0}: \mu=100$ against $H_{1}: \mu \neq 100$.


## Method 1: Using the critical value approach

$$
T=\frac{\sqrt{n}\left(\bar{X}-\mu_{0}\right)}{s}=\frac{\sqrt{100}(100.91-100)}{15.4907} \approx 0.5874
$$

The critical value is

$$
\begin{aligned}
t_{\alpha / 2 ; n-1} & =t_{0.025 ; 99} \\
& =2.000+\frac{39}{40}(1.984-2.000) \\
& =2.000+\frac{39}{40}(-0.016) \\
& =2.000-0.0156 \\
& \approx 1.984
\end{aligned}
$$

We will reject $H_{0}$ if $T \geq 1.984$ or $T \leq-1.984$ or if $|T| \geq 1.984$.
Since $-1.984<0.5874<1.984$, we do not reject $H_{0}$ at the $5 \%$ level of significance and conclude that $\mu=100$, that is, the mean IQ score is 100 .

## Method II: Using the p-value approach

$p$-value $=0.5582$. Since $0.5582>0.05$, we do not reject $H_{0}$ at the $5 \%$ level of significance and conclude that $\mu=100$, that is, the mean IQ score is 100 .
(iii) No.

We have to test $H_{0}: \sigma=15$
against $\quad H_{1}: \sigma \neq 15$

## Distributions

IQ Scores

——Normal(100.91,15.4907)

## Summary Statistics

| Mean | 100.91 |
| :--- | ---: |
| Std Dev | 15.490691 |
| Std Err Mean | 1.5490691 |
| Upper 95\% Mean | 103.98369 |
| Lower 95\% Mean | 97.836311 |
| N | 100 |

Test Standard Deviation
Hypothesized Value 15
Actual Estimate 15.4907
DF
99
Test ChiSquare
Test Statistic 105.5831
Min PValue 0.6136
Prob < ChiSq 0.6932
Prob $=$ ChiSq 0.3068

## Method 1: Using the critical value approach

Assuming $\mu$ is unknown, i.e., $\widehat{\mu}=\bar{X}$, then the test statistic is

$$
\begin{aligned}
U & =\frac{\sum_{i=1}^{n}\left(X_{i}-\bar{X}\right)^{2}}{\sigma^{2}} \\
& =\frac{(n-1) S_{X}^{2}}{\sigma^{2}} \\
& =\frac{99 \times 15.4907^{2}}{15^{2}} \\
& \approx 105.5831
\end{aligned}
$$

The critical values are

$$
\left.\left.\begin{array}{ll} 
& \chi_{1-\alpha / 2 ; n-1}^{2} \\
=\chi_{0.975 ; 99}^{2} & ;
\end{array} \begin{array}{l}
\chi_{\alpha / 2 ; n-1}^{2} \\
\chi_{0.025 ; 99}^{2}
\end{array}\right] \begin{array}{ll}
=65.6466+\frac{9}{10}(74.2219-65.6466) & ;
\end{array} \quad=118.136+\frac{9}{10}(129.561-118.136)\right)
$$

Reject $H_{0}$ if $U<73.3644$ or $U>128.4185$
Since $73.3644<105.5831<128.4185$, we do not reject $H_{0}$ at the $5 \%$ level of significance and conclude that $\sigma=15$.

## Method II: Using the p-value approach

$p$-value $=0.6136$. Since $0.6136>0.05$, we do not reject $H_{0}$ at the $5 \%$ level of significance and conclude that $\sigma=15$.
(b) We have to test: $H_{0}: \mu_{X}=\mu_{Y} \quad$ against $\quad H_{1}: \mu_{X}<\mu_{Y}$

$$
\begin{array}{lll}
n_{X}=100 & \bar{X}=100.91 & S_{X}^{2}=239.972 \\
n_{Y}=60 & \bar{Y}=105.21 & S_{Y}^{2}=240.6537
\end{array}
$$

The test statistic is

$$
T=\frac{(\bar{X}-\bar{Y})-\left(\mu_{X}-\mu_{Y}\right)}{S_{p} \sqrt{\frac{1}{n_{X}}+\frac{1}{n_{Y}}}}
$$

Now

$$
\begin{aligned}
S_{p}^{2} & =\frac{\left(n_{X}-1\right) S_{X}^{2}+\left(n_{Y}-1\right) S_{Y}^{2}}{n_{X}+n_{Y}-2} \\
& =\frac{(100-1) 239.972+(60-1) 240.6537}{100+60-2} \\
& =\frac{23757.228+14198.5683}{158} \\
& =\frac{37955.7963}{158} \\
& \approx 240.2265589 \\
& \Longrightarrow S_{\text {pooled }}=\sqrt{240.2265589} \approx 15.4992
\end{aligned}
$$

Then

$$
\begin{aligned}
T & =\frac{(\bar{X}-\bar{Y})-\left(\mu_{X}-\mu_{Y}\right)}{S_{p} \sqrt{\frac{1}{n_{X}}+\frac{1}{n_{Y}}}} \\
& =\frac{(100.91-105.21)-(0)}{15.4992 \sqrt{\frac{1}{100}+\frac{1}{60}}} \\
& =\frac{-4.3}{15.4992 \sqrt{0.026666666}} \\
& =\frac{-4.3}{2.531008761} \\
& \approx-1.6989
\end{aligned}
$$

The critical value is $t_{\alpha ;\left(n_{1}+n_{2}-2\right)}=t_{0.05 ; 158}=1.645$. Reject $H_{0}$ if $T \leq-1.645$.
Since $-1.6989<-1.645$, we reject $H_{0}$ at the $5 \%$ level of significance and conclude that $\mu_{X}<\mu_{Y}$, i.e., mean IQ score of population B is higher than the mean IQ score of population A.

## QUESTION 2

| Group | A | B | C | D |
| :---: | :---: | :---: | :---: | :---: |
| $n$ | 3 | 3 | 3 | 3 |
| $\sum X_{i j}$ | 139 | 145 | 153 | 128 |
| $\bar{X}_{i}$ | 46.3333 | 48.3333 | 51 | 42.6667 |
| $\sum\left(X_{i j}-\bar{X}_{i}\right)^{2}$ | 34.6667 | 20.6667 | 26 | 72.6667 |

(a)

$$
\begin{aligned}
S_{1}^{2} & =\frac{1}{n_{1}-1} \sum\left(X_{1 j}-\bar{X}_{2}\right)^{2} & S_{2}^{2} & =\frac{1}{n_{2}-1} \sum\left(X_{2 j}-\bar{X}_{2}\right)^{2} \\
& =\frac{1}{3-1}(34.6667) & & =\frac{1}{3-1}(20.6667) \\
& =\frac{1}{2}(34.6667) & & =\frac{1}{2}(20.6667) \\
& \approx 17.3334 & & \approx 10.3334 \\
S_{3}^{2} & =\frac{1}{n_{3}-1} \sum\left(X_{3 j}-\bar{X}_{3}\right)^{2} & S_{4}^{2} & =\frac{1}{n_{4}-1} \sum\left(X_{4 j}-\bar{X}_{4}\right)^{2} \\
& =\frac{1}{3-1}(26) & & =\frac{1}{3-1}(72.6667) \\
& =\frac{1}{2}(26) & & =\frac{1}{2}(72.6667) \\
& =13 & & \approx 36.3334
\end{aligned}
$$

From the computations above it, follows that $S_{1}^{2}=17.3334 ; S_{2}^{2}=10.3334 ; S_{3}^{2}=13$ and $S_{4}^{2}=36.3334$.
(b) (i) Ordinary average $=\frac{17.3334+10.3334+13+36.3334}{4}=\frac{77.0002}{4}=19.2501$
(ii) $M S E=\frac{S S E}{k n-k}$.

For this ANOVA problem, we have $k=4$ (there are four groups) and $n=3$ (the number of observations in each sample).

$$
\begin{aligned}
S S E & =\sum_{i=1}^{k} \sum_{j=1}^{n}\left(X_{i j}-\bar{X}_{i}\right)^{2} \\
& =34.6667+20.6667+26+72.6667 \\
& =154.0001 \\
\therefore M S E & =\frac{154.0001}{4(3)-4} \\
& =\frac{154.0001}{8} \\
& =19.2500 . \quad \text { The result in (i) = result in (ii). }
\end{aligned}
$$

This makes perfect sense! $M S E$ is like a pooled variance or an average variance, because the assumption of $A N O V A$ is that $\sigma_{1}^{2}=\sigma_{2}^{2}=\sigma_{3}^{2}=\sigma_{4}^{2}$ and if these variances are unknown, we estimate it by pooling.
(c) It is reasonable to assume that the four samples are independent. The outcome of one detergent can not influence the outcome of the other detergent.

The other assumption (of equal variances) can be formally tested!
$H_{0}: \sigma_{1}^{2}=\sigma_{2}^{2}=\sigma_{3}^{2}=\sigma_{4}^{2}$
$H_{1}: \sigma_{p}^{2} \neq \sigma_{q}^{2}$ for at least one $p \neq q$.

$$
\begin{aligned}
U & =\frac{\max S_{i}^{2}}{\min S_{i}^{2}} \\
& =\frac{36.3334}{10.3334} \\
& \approx 3.5161
\end{aligned}
$$

From Table E with $k=4$ and $v=n-1=3-1=2$, we find that the critical value is 142 . Reject $H_{0}$ if $U>142$.

Since $3.5161<142$, we cannot reject $H_{0}$ at the $5 \%$ level of significance and we may assume that the variances are equal.
(If we use the JMP computer output we may also assume that the variances are equal because "Prob > F" is not significant for all the four tests under the heading: "Tests that the Variances are Equal".)
(d) We have to test:
$H_{0}: \mu_{1}=\mu_{2}=\mu_{3}=\mu_{4}$ against
$H_{1}: \mu_{p} \neq \mu_{q}$ for at least one $p \neq q$.
The test statistic is $F=\frac{M S T_{r}}{M S E} \sim F_{k-1 ; k n-k}$

$$
\begin{aligned}
M S T_{r} & =\frac{n \sum_{i=1}^{k}\left(\bar{X}_{i}-\bar{X}\right)^{2}}{k-1} \\
\text { where } \bar{X} & =\frac{\sum \sum X_{i j}}{N}=\frac{565}{12} \approx 47.0833 \quad \text { (overall mean); } \\
\text { and } \sum\left(\bar{X}_{i}-\bar{X}\right)^{2} & =(46.3333-47.0833)^{2}+\cdots+(42.6667-47.0833)^{2} \\
& =(-0.75)^{2}+(1.25)^{2}+(3.9167)^{2}+(-4.4166)^{2} \\
& =0.5625+1.5625+15.34053889+19.50635556 \\
& \approx 36.9719 \\
\therefore M S T_{r} & =\frac{3(36.9719)}{4-1}=\frac{110.9157}{3}=36.9719
\end{aligned}
$$

We already know that $M S E=19.2500$ (see question (b)(ii)).

$$
\begin{aligned}
\therefore F & =\frac{M S T_{r}}{M S E} \\
& =\frac{36.9719}{19.25} \\
& \approx 1.9206 .
\end{aligned}
$$

(Note that these computations are the same with the JMP output under the heading: "Analysis of Variance".)

The critical value is $F_{0,05 ; 3 ; 8}=4.07$. Reject $H_{0}$ if $F>4.07$.
Since $1.9206<4.07$, we do not reject $H_{0}$ at the $5 \%$ level of significance and conclude that the population means of the four detergents do not differ, that is, $\mu_{1}=\mu_{2}=\mu_{3}=\mu_{4}$.
(Note that we reach the same conclusion with the JMP output under the heading: "Analysis of Variance" if we consider "Prob > F" $=0.2048$ )
(e) For each pair of means, we compute a test statistic

$$
T_{p q}=\frac{\bar{X}_{p}-\bar{X}_{q}}{S_{\text {pooled }} \sqrt{1 / n+1 / n}}=\frac{\sqrt{n}\left(\bar{X}_{p}-\bar{X}_{q}\right)}{\sqrt{2} S}=\frac{\sqrt{8}\left(\bar{X}_{p}-\bar{X}_{q}\right)}{\sqrt{2} \sqrt{M S E}} .
$$

We reject $H_{0}(p ; q)$ if

$$
\left|T_{p q}\right|>\sqrt{(k-1) F_{\alpha ; k-1 ; k n-k}}=\sqrt{3(4.07)} \approx 3.4943
$$

This implies that we reject $H_{0}$ if

$$
\frac{\sqrt{3}\left|\bar{X}_{p}-\bar{X}_{q}\right|}{\sqrt{2} \sqrt{0,92411}} \geq 3.4943
$$

i.e. if $\left|\bar{X}_{p}-\bar{X}_{q}\right| \geq \frac{(3.4943) \sqrt{2} \sqrt{19.25}}{\sqrt{3}}=\frac{21.68156131}{1.732050808}=12.5179$

$$
\begin{aligned}
& \left|\bar{X}_{1}-\bar{X}_{2}\right|=2<12.5179 \Longrightarrow \mu_{1}=\mu_{2} \\
& \left|\bar{X}_{1}-\bar{X}_{3}\right|=4.6667<12.5179 \Longrightarrow \mu_{1}=\mu_{3} \\
& \left|\bar{X}_{1}-\bar{X}_{4}\right|=3.6666<12.5179 \Longrightarrow \mu_{1}=\mu_{4} \\
& \left|\bar{X}_{2}-\bar{X}_{3}\right|=2.6667<12.5179 \Longrightarrow \mu_{2}=\mu_{3} \\
& \left|\bar{X}_{2}-\bar{X}_{4}\right|=5.6666<12.5179 \Longrightarrow \mu_{2}=\mu_{4} \\
& \left|\bar{X}_{3}-\bar{X}_{4}\right|=8.3333<12.5179 \Longrightarrow \mu_{3}=\mu_{4}
\end{aligned}
$$

All pairs of means are not significantly different from each other, i.e., they are equal.

## QUESTION 3

(a)

| Subject | Before | After | $Y_{i}=$ After - Before |
| :---: | :---: | :---: | :---: |
| 1 | 27 | 29 | 2 |
| 2 | 21 | 32 | 11 |
| 3 | 34 | 29 | -5 |
| 4 | 24 | 27 | 3 |
| 5 | 30 | 31 | 1 |
| 6 | 27 | 26 | -1 |
| 7 | 33 | 35 | 2 |
| 8 | 31 | 30 | -1 |
| 9 | 22 | 29 | 7 |
| 10 | 27 | 28 | 1 |
| 11 | 33 | 36 | 3 |
| 12 | 17 | 15 | -2 |
| 13 | 25 | 28 | 3 |
| 14 | 26 | 26 | 0 |
| 15 | 23 | 26 | 3 |

$$
n=15 \quad \sum Y_{i}=27 \quad \sum\left(Y_{i}-\bar{Y}\right)^{2}=198.4
$$

We have to test:
$H_{0}: \mu_{d}=0$ against
$H_{1}: \mu_{d}>0$

$$
\begin{array}{rlrl}
\bar{Y} & =\frac{1}{n} \sum Y_{i} & S_{y}^{2} & \\
& =\frac{1}{n-1} \sum\left(Y_{i}-\bar{Y}\right)^{2} \\
& =1.8 & & \\
& & & =\frac{1}{14}(198.4) \\
\Longrightarrow S_{y} & & =\sqrt{14.17142857} \\
& & & \approx 3.7645
\end{array}
$$

The test statistic is

$$
\begin{aligned}
T & =\frac{\sqrt{n}(\bar{Y}-\mu)}{S_{y}} \\
& =\frac{\sqrt{15}(1.8-0)}{3.7645} \\
& =\frac{6.971370023}{3.7645} \\
& \approx 1.8519 .
\end{aligned}
$$

The critical value is $t_{\alpha ;(n-1)}=t_{0.05 ; 14}=1.761$. Reject $H_{0}$ if $T>1.761$.
Since $1.8519>1.761$, we reject $H_{0}$ at the $5 \%$ level of significance and conclude that learning is shown by an increase in score.
(b) The output is

(c) Paired data since the two observations were taken from the same individual.

## QUESTION 4

(a) Start the JMP program
$>\quad$ Enter Engine in the first column and label it Engine.
(make sure to change the scale to nominal)
$>\quad$ Enter Reading in the second column and label it Reading.
This is a one-way ANOVA. To fit the model
$>\quad$ Choose Analyze $>$ Fit $Y$ by $X$ with Engine as $X$ factor and Reading as $Y$ response.
$>\quad$ Click Ok.
$\Longrightarrow \quad$ Then on the Oneway Analysis of Reading By Engine click on the Red triangle > Choose Unequal Variances

Oneway Analysis of Reading By Engine


Oneway Anova
Tests that the Variances are Equal


| Level | Count | Std Dev | MeanAbs to Me | MeanAbsDif to Median |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 4 | 2.645751 | 2.0000 | 2.000000 |  |
| 2 | 4 | 5.477226 | 4.0000 | - 4.000000 |  |
| 3 | 4 | 2.500000 | 1.7500 | 1.750000 |  |
| Test |  | F Ratio | DFNum | FDen | Prob $>$ F |
| O'Brien |  | 1.2740 | 2 | 9 | 0.3257 |
| Brown-F | Forsythe | 1.3774 | 2 | 9 | 0.3007 |
| Levene |  | 1.4600 | 2 | 9 | 0.2824 |
| Bartlett |  | 1.0784 | 2 |  | 0.3401 |
| Warning: Small sample sizes. Use Caution. |  |  |  |  |  |

## Welch's Test

Welch Anova testing Means Equal, allowing Std Devs Not Equal

$$
\text { F Ratio DFNum DFDen Prob }>F
$$

$$
\begin{array}{llll}
6.5121 & 2 & 5.6461 & 0.0342^{*}
\end{array}
$$

## For your own information:

The standard deviation column shows the estimates you are testing. The $p$-values are listed under the column called Prob $>F$ and are testing the assumption that the variances are equal. Small $p$-values suggest that the variance are not equal.

## Interpretation:

We have to test:
$H_{0}: \sigma_{1}^{2}=\sigma_{2}^{2}=\sigma_{3}^{2}$, against $H_{1}: \sigma_{p}^{2} \neq \sigma_{q}^{2}$ for at least one $p \neq q$
Using the Levene's test, $p$-value $=0.2824$. Since $0.2824>0.05 \Longrightarrow$ we can not reject $H_{0}$ at the $5 \%$ level of significance. The assumption of equal variances is not violated.
(b) $\Longrightarrow \quad$ Click on the triangle "Tests that the variances are equal" to hide the output.
$\Longrightarrow \quad$ Then click on the Red triangle on Oneway Analysis of Reading by Engine.
$>\quad$ Choose Means/ANOVA
$\Longrightarrow \quad$ Click again on the Red triangle and choose Means and Std dev.

Oneway Analysis of Reading By Engine


## Oneway Anova

## Summary of Fit

| Rsquare | 0.510223 |
| :--- | ---: |
| Adj Rsquare | 0.401384 |
| RootMean Square Error | 3.796929 |
| Mean of Response | 47.08333 |
| Observations (or Sum Wgts) | 12 |

Analysis of Variance

| Source | DF | Sum of <br> Squares | Mean Square | F Ratio | Prob $>$ F |
| :--- | ---: | ---: | ---: | ---: | ---: |
| Engine | 2 | 135.16667 | 67.5833 | 4.6879 | $0.0403^{*}$ |
| Error | 9 | 129.75000 | 14.4167 |  |  |
| C. Total | 11 | 264.91667 |  |  |  |

Means for Oneway Anova
Level Number Mean Std Error Lower 95\% Upper 95\%

| 1 | 4 | 45.5000 | 1.8985 | 41.205 | 49.795 |
| :--- | :--- | :--- | :--- | :--- | :--- |
| 2 | 4 | 44.0000 | 1.8985 | 39.705 | 48.295 |
| 3 | 4 | 51.7500 | 1.8985 | 47.455 | 56.045 |

3 Fid $\begin{array}{llll}4 & 51.7500 & 1.8985 & 47.455\end{array}$
Means and Std Deviations

|  |  | Std Err |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
| Level | Number | Mean | Std Dev | Mean | Lower 95\% | Upper 95\% |
| 1 | 4 | 45.5000 | 2.64575 | 1.3229 | 41.290 | 49.710 |
| 2 | 4 | 44.0000 | 5.47723 | 2.7386 | 35.285 | 52.715 |
| 3 | 4 | 51.7500 | 2.50000 | 1.2500 | 47.772 | 55.728 |

## For your information:

On the plot, the dots shows the response for each Engine. The line across the middle is the grand mean. The diamonds give a $95 \%$ confidence interval for each Engine with the middle line of each diamond showing the group mean. If the groups are significantly different, then the diamonds do not overlap.

## Interpretation:

(i) $H_{0}: \mu_{1}=\mu_{2}=\mu_{3}$ against
$H_{1}: \mu_{p} \neq \mu_{q}$ for at least one $p \neq q$.
(ii) The test statistic is $F=\frac{M S T r}{M S E} \sim F_{k-1 ; n-k}$
(iii) From the output: Computations for ANOVA we see that $F=4.6879$ which is significant with a $p$-value of 0.0403 . Since $0.0403<0.05$ we reject $H_{0}$ in favour of $H_{1}$ at the $5 \%$ level of significance and conclude that $u_{p} \neq \mu_{q}$ for at least one pair $p \neq q$, that is, the mean reading of the engines are not the same.
(c) $\Longrightarrow \quad$ Hide the output "Oneway ANOVA" and "Means and Std deviations" by clicking the triangles.
$\Longrightarrow \quad$ Click on the Red triangle on Oneway Analysis of Reading by Engine.
$\Longrightarrow \quad$ Choose Compare Means $>$ Each Pair, Student's t.

Oneway Analysis of Reading By Engine


Oneway Anova
Means Comparisons

## Comparisons for each pair using Student's t

## Confidence Quantile

| t | Alpha |
| ---: | ---: |
| 2.26216 | 0.05 |

## LSD Threshold Matrix

## Abs(Dif)-LSD

|  | 3 | 1 | 2 |
| :--- | ---: | ---: | ---: |
| 3 | -6.0735 | 0.1765 | 1.6765 |
| 1 | 0.1765 | -6.0735 | -4.5735 |
| 2 | 1.6765 | -4.5735 | -6.0735 |

Positive values show pairs of means that are significantly different.

## Connecting Letters Report

| Level |  | Mean |  |
| :--- | ---: | ---: | ---: |
| 3 | A | 51.750000 |  |
| 1 |  | B | 45.500000 |
| 2 |  | B | 44.000000 |

Levels not connected by same letter are significantly different.

## Ordered Differences Report



## Interpretation:

Yes The $A b s(D i f)-L S D$ for the pair 12 is negative. They all share the letter B and the confidence interval is ( -4.5735 ; 7.5735) and it includes zero. We conclude that the means are not significantly different from each other. Thus we conclude that $\mu_{1}=\mu_{2} \neq \mu_{3}$.
(d)

Oneway Analysis of Reading By Engine


Oneway Anova
Means Comparisons
Comparisons for all pairs using Tukey-Kramer HSD
Confidence Quantile

| q $^{*}$ | Alpha |
| ---: | ---: |
| 2.79201 | 0.05 |

LSD Threshold Matrix
Abs(Dif)-HSD

|  | 3 | 1 | 2 |
| ---: | ---: | ---: | ---: |
| 3 | -7.4961 | -1.2461 | 0.2539 |
| 1 | -1.2461 | -7.4961 | -5.9961 |
| 2 | 0.2539 | -5.9961 | -7.4961 |

Positive values show pairs of means that are significantly different.

## Connecting Letters Report

| Level |  | Mean |
| :--- | ---: | ---: |
| 3 | A | 51.750000 |
| 1 | A B | 45.500000 |
| 2 | B | 44.000000 |

Levels not connected by same letter are significantly different.

## Ordered Differences Report



Manually, we should have computed for each pair of means, a test statistic

$$
T_{p q}=\frac{\bar{X}_{p}-\bar{X}_{q}}{\mathrm{~S}_{\text {pooled }} \sqrt{\frac{1}{n}+\frac{1}{n}}}
$$

where we have samples of equal sizes if we want to incorporate the principle of the Bonferroni equality.

The Turkey-Kramer HSD that are shown in the JMP out perform individual comparisons that make adjustments for multiple test.

Confidence intervals that do not include zero imply that the pairs of means differ significantly. All pairs include zero except the pair 23. The confidence interval for the pair is ( 0.2539 : 15.2461). This is the only interval that do not include zero and it means we reject the null hypothesis and conclude that $\mu_{2} \neq \mu_{3}$. The $p$-value is 0.0431 , thus less than 0.05 and leading to the rejection of the null hypothesis of equal means.

Confirming this is the $\mathbf{A b s}$ (Dif)-LSD which is 0.2539 . Since it is positive, the means are significantly different. (Recall a negative value of Abs(Dif)-LSD means the groups are not significantly different from each other.)

## QUESTION 5

(a)

Bivariate Fit of Time By Age

(b) $H_{0}: \rho=0$

$$
\begin{aligned}
\text { against } & H_{1}: \rho>0 \\
R & =0.964
\end{aligned}
$$

The test statistic is

$$
\begin{aligned}
T & =\frac{R \sqrt{n-2}}{\sqrt{1-R^{2}}} \\
& =\frac{0.964 \sqrt{10-2}}{\sqrt{1-0.964^{2}}} \\
& =\frac{2.726603748}{\sqrt{0.070704}} \\
& =\frac{2.726603748}{0.265902237} \\
& \approx 10.2542
\end{aligned}
$$

The critical value is $t_{\alpha ;(n-2)}=t_{0.005 ; 8}=3.355$. Reject $H_{0}$ if $T \geq 3.355$.
Since $10.2542>3.355$, we reject $H_{0}$ at the $5 \%$ level of significance and conclude that $\rho>0$, that is, there is a significant positive correlation.
(c)

$$
\begin{aligned}
b & =\frac{n \Sigma a_{i} t_{i}-\left(\Sigma t_{i}\right)\left(\Sigma a_{i}\right)}{n \Sigma a_{i}^{2}-\left(\Sigma a_{i}\right)^{2}} \\
& =\frac{10(623.4)-(14.1)(394)}{10(18122)-(394)^{2}} \\
& =\frac{6234-5555.4}{181220-155236} \\
& =\frac{678.6}{25984} \\
& =0.026116071 \\
& \approx 0.0261 \\
a & =\frac{\Sigma t_{i}-b\left(\Sigma a_{i}\right)}{n} \\
& =\frac{14.1-0.026116071(394)}{10} \\
& =\frac{14.1-10.28973214}{10} \\
& =\frac{3.810267857}{10} \\
& =0.381026785 \\
& \approx 0.381
\end{aligned}
$$

The estimated regression line is $\widehat{\text { Time }}=0.381+0.0261$ Age.
(d) $a_{i}=45$

Then

$$
\begin{aligned}
\widehat{\text { Time }} & =0.381+0.0261 \text { Age } \\
& =0.381+0.0261(45) \\
& =0.381+1.1745 \\
& =1.5555 \\
& \approx 1.56 \text { seconds }
\end{aligned}
$$

$\therefore$ The predicted response time is 1.56 seconds.
(e) $\widehat{\beta}_{1}=0.0261$

For every increase in age of 1 year, the response time increase by 0.0261 seconds.
(f) The standard error of the estimate is given by $S E=S \sqrt{\frac{1}{n}+\frac{\left(a_{i}-\bar{a}\right)^{2}}{d^{2}}}$.

| $X$ | $Y$ | $\widehat{\beta}_{0}+\widehat{\beta}_{1} X$ | $e_{i}=Y_{i}-\left(\widehat{\beta}_{0}+\widehat{\beta}_{1} X\right)$ | $e_{i}^{2}$ |
| :--- | :--- | ---: | ---: | :--- |
| 20 | 0.9 | 0.903348 | -0.003348 | 0.000011 |
| 23 | 0.8 | 0.981696 | -0.181696 | 0.033013 |
| 25 | 1.3 | 1.033929 | 0.266071 | 0.070794 |
| 31 | 1.1 | 1.190625 | -0.090625 | 0.008213 |
| 33 | 1.2 | 1.242857 | -0.042857 | 0.001837 |
| 38 | 1.5 | 1.373438 | 0.126562 | 0.016018 |
| 40 | 1.4 | 1.425670 | -0.02567 | 0.000659 |
| 53 | 1.7 | 1.765179 | -0.065179 | 0.004248 |
| 59 | 1.9 | 1.921875 | -0.021875 | 0.000479 |
| 72 | 2.3 | 2.261384 | 0.038616 | 0.001491 |
|  |  |  |  | 0.136763 |

$$
\begin{aligned}
M S E & =s^{2} \\
& =\frac{1}{n-2} \sum\left(Y_{i}-\widehat{\beta}_{0}-\widehat{\beta}_{1} X\right)^{2} \\
& =\frac{1}{10-2}(0.136763) \\
& =\frac{1}{8}(0.136763) \\
& =0.017095375 \\
\Longrightarrow s & =\sqrt{0.017095375} \\
& \approx 0.1307
\end{aligned}
$$

$$
\begin{aligned}
d^{2} & =\sum\left(X_{i}-\bar{X}\right)^{2} \\
& =\Sigma a_{i}^{2}-\frac{\left(\Sigma a_{i}\right)^{2}}{n} \\
& =18122-\frac{(394)^{2}}{10} \\
& =2598.4 \\
\bar{a}_{i} & =\frac{394}{10} \\
& =39.4
\end{aligned}
$$

Then

$$
\begin{aligned}
S E & =S \sqrt{\frac{1}{n}+\frac{\left(a_{i}-\bar{a}\right)^{2}}{d^{2}}} \\
& =0.1307 \sqrt{\frac{1}{10}+\frac{(45-39.4)^{2}}{2598.4}} \\
& =0.1307 \sqrt{0.1+\frac{31.36}{2598.4}} \\
& =0.1307 \sqrt{0.1+0.012068965} \\
& =0.1307 \sqrt{0.112068965} \\
& \approx 0.0438
\end{aligned}
$$

(g) The confidence interval for $\beta_{1}$ is

$$
\begin{array}{ll} 
& \widehat{\beta}_{1} \pm t_{\alpha / 2 ; n-2} \times \frac{s}{d} \\
\widehat{\beta}_{1}= & t_{\alpha / 2 ; n-2}=t_{0.025 ; 8}=2.306 \\
d=\sqrt{2598.4} \approx 50.9745 & s=0.1307
\end{array}
$$

Thus, the $95 \%$ confidence interval for the slope $\widehat{\beta}_{1}$ is 0.0261

| $\widehat{\beta}_{1}$ | $\pm$ | $t_{\alpha / 2 ; n-2} \times \frac{s}{d}$ |
| :--- | :--- | :--- |
| 0.0261 | $\pm$ | $2.306 \times \frac{0.1307}{50.9745}$ |
| 0.0261 | $\pm$ | $2.306 \times 0.002564027$ |
| 0.0261 | $\pm$ | 0.0059 |
| $(0.0261-0.0059$ | $;$ | $0.0261+0.0059)$ |
| $(0.0202$ | $;$ | $0.032)$ |

(h) The JMP output is

Bivariate Fit of Time By Age


## Linear Fit

Time $=0.3810268+0.0261161^{\star}$ Age
Summary of Fit
RSquare $\quad 0.928359$
RSquare Adj 0.919403

RootMean Square Error $\quad 0.130749$
Mean of Response 1.41
Observations (or Sum Wgts) 10

## Analysis of Variance

|  |  | Sum of <br> Squares | Mean Square | F Ratio |
| :--- | ---: | ---: | ---: | ---: |
| Source | DF | 1.7722366 | 1.77224 | 103.6673 |
| Model | 1 | 0.01710 | Prob $>$ F |  |
| Error | 8 | 0.1367634 | 0.0170 |  |
| C. Total | 9 | 1.9090000 |  | $<.0001^{*}$ |

## Parameter Estimates

| Term | Estimate | Std Error | t Ratio | Prob> $\|\mathrm{t}\|$ |
| :--- | ---: | :--- | :--- | ---: | :--- |
| Intercept | 0.3810268 | 0.109192 | 3.49 | $0.0082^{*}$ |
| Age | 0.0261161 | 0.002565 | 10.18 | $<.0001^{*}$ |


| Correlation |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: |
| Variable | Mean | Std Dev Correlation | Signif. Prob | Number |  |
| Age | 39.4 | 16.9915 | 0.963514 | $<.0001^{*}$ | 10 |
| Time | 1.41 | 0.460555 |  |  |  |

